



# Financial Risk Management and Governance **Volatilities**

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## Definition

- The volatility of a variable is the standard deviation of its return with the return being expressed with continuous compounding
- The variance rate is the square of volatility
- Implied volatilities are the volatilities implied from option prices
- Normally days when markets are closed are ignored in volatility calculations (252 days)

 $\widehat{\sigma\sqrt{T}}$  standard deviation of  $\ln(S(T)/S(0))$ 



When T is small

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- » The continuously compounded return of a market variable is close to the percentage change
- »  $\sigma\sqrt{T}$   $\approx$  std. deviation of % change in market variable



Some thoughts on implied volatilities...

current estimation o reglied volats of the marker 1 + the desire" to be protected = the "how much 3 lestrical whate  $=$   $\sqrt{20}$  arriver. Ca ve me a (finish tune univeloir the rolat. we would loke to make the hostswical volat more "voighted" to recent observations





# The VIX index Inglish Uslat, Inder





#### From historical data...

#### ■ We have

- » *n+1* price observations
- » Returns: i=1, 2..., n

$$
u_i = \ln\left(\frac{S_i}{S_{i-1}}\right)
$$

» The estimate of the standard deviation of  $u_i$  is given by

$$
s = \sqrt{\frac{1}{n-1} \sum_{i=1}^{n} (u_i - \overline{u})^2}
$$

or

or  

$$
s = \sqrt{\frac{1}{n-1} \sum_{i=1}^{n} u_i^2 - \frac{1}{n(n-1)} (\sum_{i=1}^{n} u_i)^2}
$$

 $\rightarrow$  s is an estimate of  $\sigma\sqrt{T}$ 

**b** Therefore: 
$$
\hat{\sigma} = \frac{s}{\sqrt{T}}
$$
 with standard error  $\approx \hat{\sigma}/\sqrt{2n}$ 



#### Are Daily Changes in Exchange Rates Normally Distributed?

#### Knowing this...



 What should you do if you observe fat tails in the tails of a market variable while the market assumes log-normality of prices?





### The power law



H. Pirotte **8**

#### Formulation

**»** For many variables, when x is large (*K* and  $\alpha$  are constants)

$$
Pr(v > x) = Kx^{-\alpha}
$$

- This seems to fit the behavior of the returns on many market variables better than the normal distribution This seems to fit the behavior of the ret<br>variables better than the normal distrib<br>The previous equation implies that<br> $\ln [Pr(v > x)] = \ln K -$ <br>This law will be used when we will be ta<br>Value Theory and Operational Risk.
- The previous equation implies that

$$
\ln[\Pr(v > x)] = \ln K - \alpha \ln x
$$

This law will be used when we will be talking about Extreme





## Standard Approach to Estimating Volatility

- **•** Define  $\sigma_n$  as the volatility per day between day  $n$ -1 and day  $n$ , as estimated at end of day *n-*1
- Define  $S<sub>i</sub>$  as the value of market variable at end of day *i*
- **Define**  $u_i = \ln(S_i/S_{i-1})$

$$
\sigma_n^2 = \frac{1}{m-1} \sum_{i=1}^m (u_{n-i} - \overline{u})^2
$$

$$
\overline{u} = \frac{1}{m} \sum_{i=1}^m u_{n-i}
$$

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Simplifications Usually Made in Risk Management

- Define  $u_i$  as  $(S_i-S_{i-1})/S_{i-1}$
- **Assume that the mean value of**  $u_i$  **is zero**
- Replace *m-*1 by *m*

This gives

$$
\sigma_n^2 = \frac{1}{m} \sum_{i=1}^m u_{n-i}^2
$$



#### Weighting Scheme

Instead of assigning equal weights to the observations we can set





#### ARCH(m) model

I In an ARCH(m) model we also assign some weight to the long-run variance rate, *V<sup>L</sup>* :

2 2 *V u* 1 1 where 1 *m n L i n i <sup>i</sup> m i i* 

$$
AR(U(L))
$$
:  $\sigma_{m}^{2} = \gamma V_{L} + (1-\gamma) W_{m-1}$ 

#### EWMA

 In an exponentially weighted moving average model, the weights assigned to the  $u^2$  decline exponentially as we move back through time  $\overbrace{\sigma_n^2 = \omega \sigma_{n-1}^2 + (1-\lambda)u_{n-1}^2}^{\text{max}}$  $-\lambda \frac{\sqrt{1-\lambda^2}}{2a_{n-1}^2+\lambda^2}$ 

 $\frac{1}{1} + (1 - \lambda) u_{n-1}^2$ 

 $\frac{1}{1-(1-\lambda)u_{n-1}^2}$ 

*u*

 $\widehat{(1 - \lambda)}$ 

 $-\frac{\mu_{n-1}}{2}$ <br>=  $(1-\lambda)\sum_{n=1}^{m} \lambda^{i-1} u_{n-i}^2 + \lambda^m \sigma_{n-m}^2$  $\lambda$ )<br>  $\sum_{n=1}^{m} \lambda^{i-1} u_{n-i}^2 + \lambda^m \sigma_{n-m}^2$  $\sum_{n=1}^{\infty}$ <br>(1 -  $\lambda$ )  $\sum_{m=1}^{m} \lambda^{i-1} u^2 + \lambda^m \sigma^2$ *i*  $=$ 1 the residual  $\sigma^2$  for in the past. لحماها Advantages » Relatively little data needs<sup>t</sup> to be stored

 $\frac{d^{2}}{2} = \lambda \sigma^{2} + (1 - \lambda)u^{2}$ 

 $\int_{n}^{2} = \left(\lambda \sigma_{n-1}^{2} + \left(1 - \lambda\right)u_{n}^{2}\right)$ 

- » We need only remember the current estimate of the variance rate and the most recent observation on the market variable
- » Tracks volatility changes

This leads to

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**»** RiskMetrics uses  $\lambda$  = 0.94 for daily volatility forecasting





» The long-run variance rate is 0.0002 so that



## GARCH (1,1)

- Example 2
	- **»** Suppose that the current estimate of the volatility  $\left(s\right)$  1.6% per day and the most recent percentage change in the market variable is 1%
	- » The new variance rate is

 $0.000002 + 0.13 \times 0.0001 + 0.86 \times 0.000256 + 0.00023336$ 

The new volatility is 1.53% per day



#### GARCH (p,q)  $\sigma_n^2 = \omega + \sum \alpha_i u_{n-i}^2 + \sum \beta_j \sigma_i$ *j*  $\sum_{i=1}^{p} \alpha_i u_{n-i}^2 + \sum_{j=1}^{q}$  $\frac{2}{n} = \omega + \sum \alpha_i u_{n-i}^2 + \sum \beta_j \sigma_{n-j}^2$ 1 1  $\alpha=\omega+\sum\alpha_{i}u_{n-i}^{2}+\sum\beta_{j}\sigma_{n-i}^{2}$  $=1$   $j=$  $\sum \alpha_i^{} u_{n-i}^2 + \sum \beta_j^{} \sigma_{n-i}^2$

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## **Others**

- **We can design GARCH models so that the weight given to**  $u_i^2$ depends on whether  $u_i$  is positive or negative
- We do not have to assume that the conditional distribution is normal



#### Variance targeting

- One way of implementing GARCH(1,1) that increases stability is by using variance targeting
- We set the long-run average volatility equal to the sample variance
- Only two other parameters then have to be estimated



## Maximum likelihood

- In maximum likelihood methods we choose parameters that maximize the likelihood of the observations occurring
- Example 1
	- » We observe that a certain event happens one time in ten trials. What is our estimate of the proportion of the time, *p*, that it happens?
	- **»** The probability of the outcome is  $p(1-p)^9$
	- » We maximize this to obtain a maximum likelihood estimate: *p=*0.1 (differentiate with respect to *p* and set result equal to 0)
- Example 2
	- » Estimate the variance of observations from a normal distribution with mean zero 1  $\int -u_i^2$ Maximize:  $\left| \right| \left| \frac{1}{\sqrt{2}} \exp \left| \frac{1}{\sqrt{2}} \right| \right|$ *m*  $u_i^{\prime}$ rvations from a normal dis<br> $\left[\frac{1}{\sqrt{2\pi}}\exp\left(\frac{-u_i^2}{\sqrt{2\pi}}\right)\right]$ rvations from a normal dis<br> $\left[\frac{1}{\sqrt{2\pi v}}\exp\left(\frac{-u_i^2}{2v}\right)\right]$  $\prod$

Maximize **c. EXECUTE:** 
$$
\prod_{i=1}^{m} \left[ \frac{1}{\sqrt{2\pi v}} \exp\left(\frac{-u_i^2}{2v}\right) \right]
$$
  
or: 
$$
\sum_{i=1}^{m} \left[ -\ln(v) - \frac{u_i^2}{v} \right] = -m \ln(v) - \sum_{i=1}^{m} \left[ \frac{u_i^2}{v} \right]
$$
  
This gives: 
$$
v = \frac{1}{m} \sum_{i=1}^{m} u_i^2
$$

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### Application to GARCH(1,1)

We choose parameters that maximize

